Exercise session 4

- 1. Your aim is to estimate how the number of prenatal examinations and several other characteristics influence the birth weight of a baby. Your initial hypothesis is that more responsible pregnant women visit the doctor more often and this leads to healthier and thus also bigger babies.
- (a) In your first specification, you run the following model:

$$bwght = \beta_0 + \beta_1 npvis + \beta_2 npvis^2 + \beta_3 monpre + \beta_4 male + \varepsilon$$
,

where *bwght* is birth weight of the baby (in grams), *npvis* is the number of prenatal doctor's visits, *monpre* is the month on pregnancy in which the prenatal care began and *male* is a dummy, equal to one if the baby is a boy and zero if it is a girl. You obtain the following results from Stata¹:

Source	SS	df	MS		Number of obs	
Model Residual	12848047.5 570003184		12011.87 1204.639		Prob > F R-squared	= 9.70 $= 0.0000$ $= 0.0220$ $= 0.0100$
Total	582851231	1725 33	7884.772		Adj R-squared Root MSE	= 0.0198 = 575.5
bwght	Coef.	Std. Err	. t	P> t	[95% Conf.	INTERVAL]
npvis npvissq monpre MALE _CONS	53.50974 -1.173175 30.47033 76.69243 2853.196	11.41313 .3591552 12.40794 27.76083 101.3073	-3.27 2.46 2.76	0.000 0.001 0.014 0.006 0.000	31.12468 -1.877601 6.134091 22.24391 2654.498	75.8948 4687481 54.80657 131.141 3051.895

- i. Is there strong evidence that *npvissq* (stands for *npvis*²) should be included in the model? The *p*-value on the coefficient on *npvissq* is very small, and hence the variable is strongly significant and should be included in the model.
- ii. How do you interpret the negative coefficient of npvissq? The negative coefficient on npvissq signals a concave form of the impact of the number of prenatal doctor's visits, meaning that there are decreasing returns to visiting the doctor. A possible explanation is that some number of visits is beneficiary for all pregnant women, but higher necessity of visits could

¹ Stata is a statistical software, which can be used to for econometric purposes. The Stata output

is quite similar to the Gretl output you are familiar with. In particular, *Coef.* denotes the estimated coefficients, *Std.Err.* denotes the standard errors of these coefficients, *t* denotes the *t*-statistic of the test of significance of the coefficients, P > |t| denotes the corresponding *p*-value.

mean that the pregnancy is risky for some reasons and the woman has to go to the doctor more often than usually. Such woman is also more likely to have smaller baby.

iii. Holding *npvis* and *monpre* fixed, test the hypothesis that newborn boys weight by 100 grams more than newborn girls (at 95% confidence level).

Such hypothesis can be stated as

$$H_0: \beta_4 = 100 \ H_a: \beta_4 \neq 100$$

Test statistic $t = \frac{\widehat{\beta_4} - 100}{SE(\widehat{\beta_4})} = \frac{76.69 - 100}{27.76} = -0.84 \sim t_{\infty,1721}$ =-1.96. Therefore, we failed to reject the null hypothesis that newborn boys weigh by 100 grams more than newborn girls at 95% confidence level.

b. A friend of yours, student of medicine, reminds you of the fact that the age of the parents (especially of the mother) might be a decisive factor for the health and for the weight of the baby. Therefore, in your second specification, you decide to include in your model also the age of the mother (*mage*) and of the father (*fage*). The results of your estimation are now the following:

Source	SS	df	MS			Number of obs	
Model Residual	16270165.8 563258231	6 1713	2711694 328813.9			R-SQUARED	= 0.0000 = 0.0281
Total	579528396	1719	337131.1	21			= 0.0247 = 573.42
bwght	Coef.	Std.	Err.	t	P> t	[95% Conf.	INTERVAL]
npvis npvissq monpre MALE MAGE FAGE _cons	52.43859 -1.138545 34.35661 74.45482 .5285275 8.697342 2592.813	11.40 .3585 12.69 27.75 4.218 3.465 139.6	648 - 477 247 069 973	4.60 3.18 2.71 2.68 0.13 2.51 8.57	0.000 0.002 0.007 0.007 0.900 0.012 0.000	30.06826 -1.841816 9.457725 20.02252 -7.744582 1.899357 2318.974	74.80891 4352743 59.2555 128.8871 8.801637 15.49533 2866.651

i. Comment on the significance of the coefficients on *mage* and *fage* separately: are they in line with your friend's claim?

When we look on the *p*-values of the corresponding coefficients, we see that whereas *fage* is significant at 99% confidence level, *mage* is insignificant. This is not in line with our friend's claim, who says that especially the age of the mother should be an important factor.

ii. Test the hypothesis that mage and fage are jointly significant (at 95% confidence level). Is the result in line with your friend's claim? To test joint significance, we need restricted and unrestricted models. In the regression in part (b) we have included mage and fage while they are not included in the regression in part (a). Therefore, we can use SSR from both regression outputs in order to judge the

joint significance of the *mage* and *fage* variables. According to output in part (a) SSR_r=570003184, According to output in part (b) SSR_{ur}=563258231. We construct F test based on the formula: $F = \frac{(SSR_r - SSR_{ur})/q}{SSR_{ur}/df}$, where q is the number of restrictions in this case q=2 (mage and fage) and df is degrees of freedom. Df=n-k-1=1720-7 Therefore $F = \frac{(570003184 - 563258231)/2}{2} = 10,36$ in the E-table we will find a

Therefore, $F = \frac{(570003184 - 563258231)/2}{563258231/1713} = 10.36$ in the F-table we will find a critical value at 5% it will be $F_{2,\infty} = 3.00$.

10.36>3, hence, we can reject the null hypothesis and we conclude that mage and fage are jointly significant.

iii. How can you reconcile you findings from the two previous questions?

The finding about the joint significance from the second question is not surprising, since we know already from the first question that fage is individually significant. If a variable is significant, then the H_A of the test of the joint significance has to be valid and so the variables have to be jointly significant.

c) In your third specification, you decide to drop fage and you get the following results:

Source	SS	df	MS		Number of obs	
Model Residual	14451685.6 568399545		890337.13 30464.852		F(5, 1720) = Prob > F R-squared	= 0.0000 = 0.0248
Total	582851231	1725 3	37884.772		Adj R-squared Root MSE	= 0.0220 = 574.86
bwght	Coef.	Std. E	Crr. t	P> t	[95% Conf	. INTERVAL]
npvis npvissq monpre MALE MAGE _cons	52.27885 -1.142647 35.25912 79.38175 -6.91257 2648.851	11.414 .35902 12.583 27.756 3.1379 137.27	214 -3.18 328 2.80 567 2.86 972 -2.20	0.000 0.001 0.005 0.004 0.028 0.000	29.89196 -1.846811 10.57898 24.94136 -13.06721 2379.602	74.66575 4384821 59.93927 133.8221 757928 2918.1

Comment on the significance of the coefficient on *mage*, compared to the results from part (b). Is your finding in line with your reasoning in part (b)? Does it confirm your friend's claim?

Now, the *p*-value of the coefficient on *mage* is very low and so the coefficient is strongly significant. When we compare this finding to part (b), we realize that the insignificance of this coefficient in that part was probably given by a strong correlation between *mage* and *fage*, leading to the multicollinearity problem, which increases the standard errors and decreases thus the significance of the coefficients. When we drop *fage*, the multicollinearity problem is solved and we see that our friend's claim was true.

d) Having regained trust in your friend, you consult your results once more with him. Together, you come up with an interesting question: whether smoking during pregnancy can affect the weight of the baby. Fortunately, you have at your disposition the variable *cigs*, standing for the average number of cigarettes each woman in your sample smokes per day during the pregnancy, and so you can include it in your model. However, your friend warns you that women who smoke during pregnancy are in general less responsible than those who do not smoke, and that these women also tend to visit the doctor less often. (In other words, the more the women smokes, the less prenatal doctor's visits she has). This is an important fact that you have to take into consideration while interpreting your final results, which are:

Source Model RESIDUAL	SS 14560828.9 523281374		MS 6804.81 013.235		Number of obs F(6, 1615) Prob > F R-squared	
Total	537842203	1621 331	796.547		Adj R-squaren Root MSE	= 0.0235 = 569.22
bwght	Coef.	Std. Err.	t	P> t	[95% Conf	. INTERVAL]
npviss npvissq monpre MALE MAGE cigs _cons	42.43442 8948737 31.77658 82.39438 -6.980738 -10.209 2748.856	11.59582 .3624432 12.78156 28.34937 3.227181 3.398309 141.868	3.66 -2.47 2.49 2.91 -2.16 -3.00 19.38	0.000 0.014 0.013 0.004 0.031 0.003 0.000	19.68999 -1.605782 6.706395 26.78897 -13.31064 -16.87456 2470.591	65.17885 1839653 56.84676 137.9998 6508356 -3.54344 3027.12

i. Interpret the coefficient on *cigs*.

The coefficient on *cigs* tells us that with each additional cigarette smoked by the pregnant woman on average per day, the weight of the baby is smaller by 10 grams, ceteris paribus.

ii. What evidence do you find that *cigs* really should be included in the model? List at least two arguments.

We can see from the *p*-value that the coefficient on cigs is strongly significant. We can also see that the R^2 as well as the adjusted R^2 are higher than in the model without this variable (in part (c)). Moreover, we see that the coefficient on *npvis* has changed quite a lot once we included *cigs*, which is a signal of an omitted variable bias in part (c) and a proof that *cigs* indeed should be included in the model.

iii. Compare the coefficient on *npvis* with the one you obtained in part (c). Do you think there was a bias? If yes, explain where it came from and interpret its sign.

In part (c), the coefficient on npvis was approximatively equal to 52, now it is equal to 42. This shows there was a positive bias in part (c): the coefficient was

overestimated there. We know that the sign of this bias is the sign of the product of two correlations: the correlation between the omitted variable *cigs* and the variable *npvis* and the correlation between *cigs* and the dependent variable *bwght*. The correlation between *cigs* and the dependent variable *bwght* is negative as we can see from the negative coefficient on *cigs* in the model estimated in part (d), the correlation between *cigs* and *npvis* is negative as we learn from our friend (women who smoke tend to visit the doctor less often). The product of these two correlations is thus positive and so is the bias in part (c).

Intuitively, we can say that when *cigs* was omitted, everything that could measure the degree of responsibility of pregnant women in our model was the variable *npvis*. Once we included *cigs*, we can measure separately the responsibility of going to the doctor and the responsibility of not smoking, and so the coefficient on *npvs* is reflecting only the correct part of this influence and it is not overestimated.