

# LECTURE 3

Introduction to Econometrics

## INTRODUCTION TO LINEAR REGRESSION ANALYSIS II

October 6, 2017

# REVISION: THE PREVIOUS LECTURE

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  - ▶ An estimator is **efficient** if the variance of its sampling distribution is the smallest possible

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- ▶ We found the formulae for the estimates:

$$\hat{\beta}_1 = \frac{\sum_{i=1}^n (x_i - \bar{x}_n) (y_i - \bar{y}_n)}{\sum_{i=1}^n (x_i - \bar{x}_n)^2} \quad \hat{\beta}_0 = \bar{y}_n - \hat{\beta}_1 \bar{x}_n$$



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- ▶ Remember that all of these factors are included in the error term and may alter its properties
- ▶ The properties of the error term determine the properties of the estimates

## WARM-UP EXERCISE

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1. Interpret the meaning of the coefficient of  $exper_i$ .
2. Use the estimates to determine the average wage of a person with 1, 5, 20, and 40 years of experience.
3. Do the predicted wages seem realistic? Explain your answer.

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- ▶ Readings:
  - ▶ Studenmund - chapter 4
  - ▶ Wooldridge - chapters 5, 8, 9, 12



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- ▶ For observations  $1, 2, \dots, n$ , we have:

$$\begin{aligned}y_1 &= \beta_0 + \beta_1 x_{11} + \beta_2 x_{12} + \dots + \beta_k x_{1k} + \varepsilon_1 \\y_2 &= \beta_0 + \beta_1 x_{21} + \beta_2 x_{22} + \dots + \beta_k x_{2k} + \varepsilon_2 \\&\vdots \\y_n &= \beta_0 + \beta_1 x_{n1} + \beta_2 x_{n2} + \dots + \beta_k x_{nk} + \varepsilon_n\end{aligned}$$

# MATRIX NOTATION

- ▶ We can write in matrix form:

$$\begin{pmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{pmatrix} = \begin{pmatrix} 1 & x_{11} & x_{12} & \cdots & x_{1n} \\ 1 & x_{21} & x_{22} & \cdots & x_{2n} \\ \vdots & \vdots & \vdots & & \vdots \\ 1 & x_{n1} & x_{n2} & \cdots & x_{nk} \end{pmatrix} \begin{pmatrix} \beta_0 \\ \beta_1 \\ \beta_2 \\ \vdots \\ \beta_k \end{pmatrix} + \begin{pmatrix} \varepsilon_1 \\ \varepsilon_2 \\ \vdots \\ \varepsilon_n \end{pmatrix}$$

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or in a simplified notation:

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon}$$

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$$\begin{aligned}\hat{\beta} &= \operatorname{argmin}_{\beta} (\mathbf{y} - \mathbf{X}\beta)' (\mathbf{y} - \mathbf{X}\beta) \\ &= \operatorname{argmin}_{\beta} \mathbf{y}'\mathbf{y} - \mathbf{y}'\mathbf{X}\beta - \beta'\mathbf{X}'\mathbf{y} + \beta'\mathbf{X}'\mathbf{X}\beta\end{aligned}$$

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- ▶ This gives us

$$\hat{\beta} = (\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}'\mathbf{y}$$

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$$Q = \beta_0 + \beta_1 P + \beta_2 P_s + \beta_3 Y + \varepsilon$$

estimated as  $\hat{Q} = 31.50 - 0.73P + 0.11P_s + 0.23Y$

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- ▶ Meaning of  $\beta_1$  is the impact of a one unit increase in  $P$  on the dependent variable  $Q$ , **holding constant the other included independent variables**  $P_s$  and  $Y$
- ▶ When price increases by 1 unit (and price of substitute good and income remain the same), quantity demanded decreases by 0.73 units

## EXERCISE

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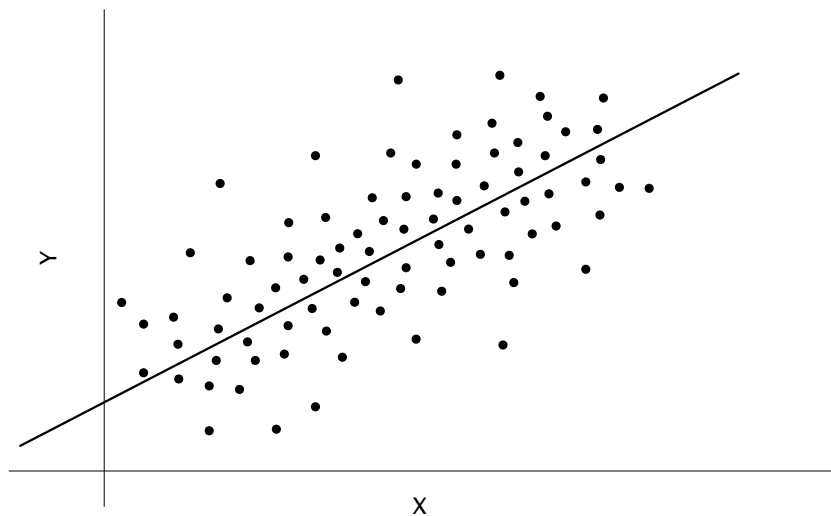
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7. The error term is normally distributed

# GRAPHICAL REPRESENTATION



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- ▶ Note that it is the linearity in coefficients that allows us to rewrite the general regression model in matrix form

## EXERCISE

Which of the following models is/are linear?

▶  $y = \beta_0 + \beta_1 x + \varepsilon$

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- ▶  $y = x^{\beta_1} + \varepsilon$  is NOT a linear model
- ▶ Regression models are **linear in parameters**, but they do not need to be linear in variables

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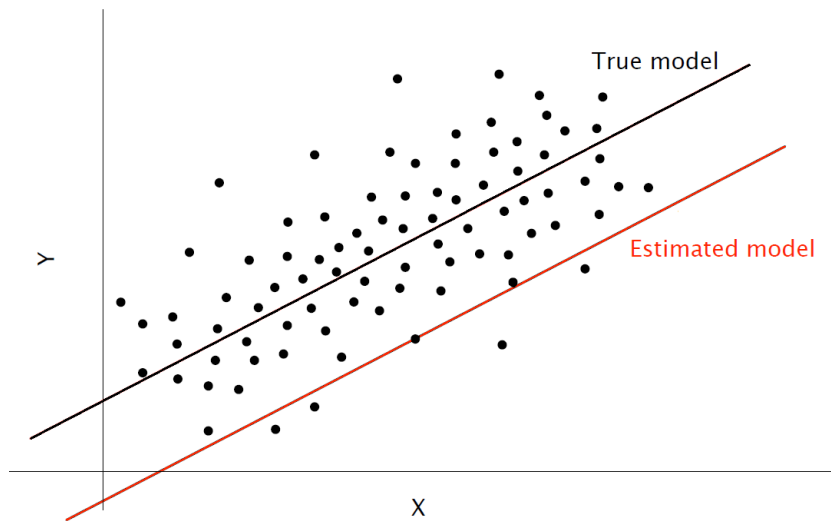
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- ▶ In fact, the mean of  $\varepsilon_i$  is forced to be zero by the existence of the intercept ( $\beta_0$ ) in the equation
- ▶ Hence, this assumption is satisfied as long as there is an **intercept** included in the equation

# GRAPHICAL REPRESENTATION





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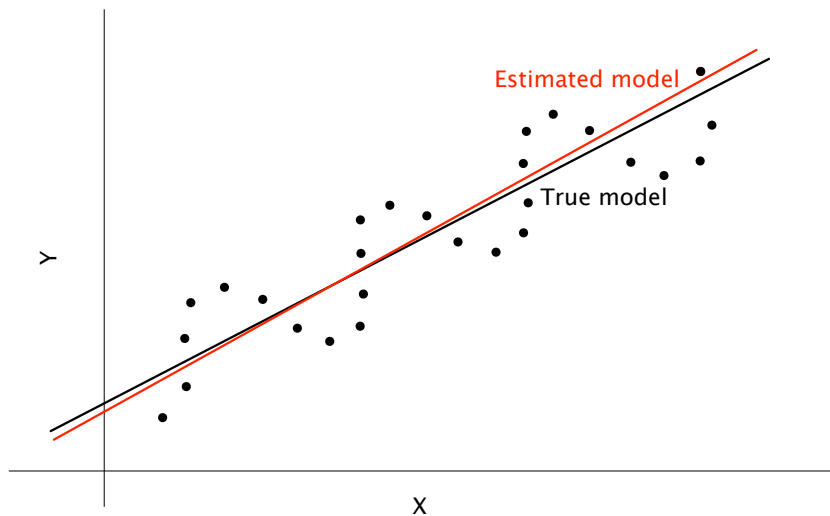
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- ▶ We will solve this problem using Generalized Least Squares estimator

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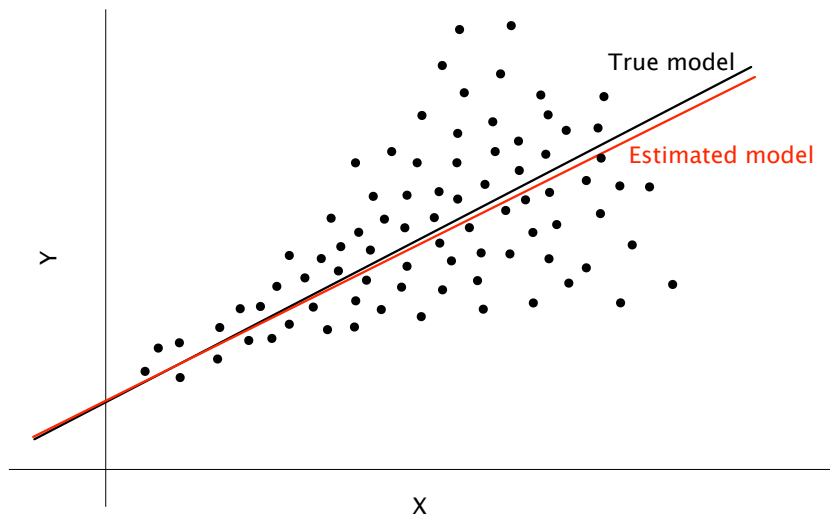
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- ▶ We will solve this problem using Hull-White robust standard errors



# GRAPHICAL REPRESENTATION



### 3. NO CORRELATION + 4. HOMOSKEDASTICITY

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► Notation:

- no correlation:  $\text{corr}(\varepsilon_i \varepsilon_j) \Rightarrow E[\varepsilon_i \varepsilon_j] = 0$  for each  $i, j$
- homoskedasticity:  $E[\varepsilon_i^2] = \sigma^2$  for each  $i$

► Matrix notation:

$$\text{Var}[\varepsilon] = \begin{pmatrix} \sigma^2 & 0 & 0 & \dots & 0 \\ 0 & \sigma^2 & 0 & \dots & 0 \\ 0 & 0 & \sigma^2 & \dots & 0 \\ & \vdots & & \ddots & \vdots \\ 0 & 0 & 0 & \dots & \sigma^2 \end{pmatrix} = \sigma^2 \mathbf{I}$$

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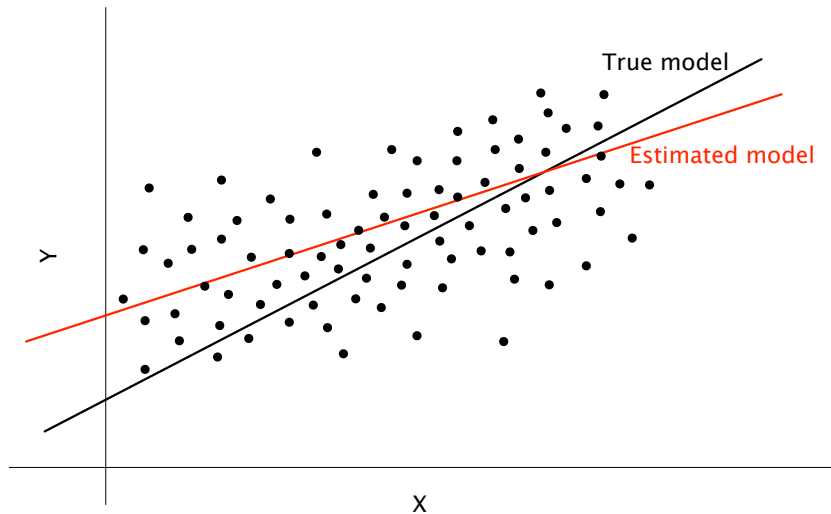
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- ▶ We will solve this problem using IV approach

# GRAPHICAL REPRESENTATION



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  - ▶ Example: we include dummy variables for men and women together with the intercept

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- ▶ Knowing the distribution of the estimate allows us to find its confidence intervals and to test hypotheses about coefficients

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- ▶ The properties of  $\hat{\beta}$  are based on the properties of  $\varepsilon$

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- ▶ Note that the normality of errors is not required for large samples, because  $\hat{\beta}$  is asymptotically normal anyway

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- ▶ As a consequence, the assumption of zero correlation between explanatory variables and the error term

$$E[\mathbf{X}'\varepsilon] = \mathbf{0}$$

is the most important one to satisfy in regression models

## SUMMARY

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  - ▶ (normally distributed) error term with zero mean and constant variance, no serial autocorrelation
  - ▶ no correlation between error term and explanatory variables
- ▶ We showed that if these assumptions hold, OLS estimate is
  - ▶ consistent (if no correlation between  $\mathbf{X}$  and  $\boldsymbol{\varepsilon}$ )
  - ▶ unbiased (if no correlation between  $\mathbf{X}$  and  $\boldsymbol{\varepsilon}$ )
  - ▶ efficient (if homoskedasticity and no autocorrelation of  $\boldsymbol{\varepsilon}$ )

## SUMMARY

- ▶ We expressed the multivariate OLS model in matrix notation  $\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon}$  and we found the formula of the estimate:  
$$\hat{\boldsymbol{\beta}} = (\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}'\mathbf{y}$$
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  - ▶ efficient (if homoskedasticity and no autocorrelation of  $\boldsymbol{\varepsilon}$ )
  - ▶ normally distributed (if  $\boldsymbol{\varepsilon}$  normally distributed)