#### Econometrics - Lecture 1

## Econometrics – First Steps

#### Contents

- Organizational Issues
- Some History of Econometrics
- An Introduction to Linear Regression
  - OLS: An Algebraic Tool
  - The Linear Regression Model
  - Small Sample Properties of the OLS Estimator
- Introduction to GRETL

### Organizational Issues

#### **Course schedule**

Class	Date	
1	Fr, Sept 29	
2	Fr, Oct 6	
3	Fr, Oct 13	
4	Fr, Oct 27	
5	Fr, Nov 3	
6	Fr, Nov 10	

#### Time: 10:00-14:00

#### Aims of the course

- Use of econometric tools for analyzing economic data: specification of adequate models, identification of appropriate econometric methods, estimation of model parameters, interpretation of results
- Introduction to commonly used econometric tools and techniques
- Understanding of econometric concepts and principles
- Use of GRETL

### Example: Individual Wages

Sample (US National Longitudinal Survey, 1987)

- N = 3294 individuals (1569 females)
- Variable list
  - WAGE: wage (in 1980 \$) per hour (p.h.)
  - MALE: gender (1 if male, 0 otherwise)
  - SCHOOL: years of schooling
  - EXPER: experience in years
  - AGE: age in years
- Questions of interest
  - Effect of gender on wage p.h.: Average wage p.h.: 6,31\$ for males, 5,15\$ for females
  - □ Effects of education, of experience, of interactions, etc. on wage p.h.

# Example: Income and Consumption



#### Literature

Course textbook

- Marno Verbeek, A Guide to Modern Econometrics, 4<sup>rd</sup> ed., Wiley, 2012
- Suggestions for further reading
- Peter Kennedy, A guide to econometrics. 6th ed., Blackwell, 2008.
- William H. Greene, *Econometric Analysis*. 7th Ed., Prentice Hall, 2011

#### **Prerequisites**

- Linear algebra: linear equations, matrices, vectors (basic operations and properties)
- Descriptive statistics: measures of central tendency, measures of dispersion, measures of association, frequency tables, histogram, scatter plot, quantile
- Theory of probability: probability and its properties, random variables and distribution functions in one and in several dimensions, moments, convergence of random variables, limit theorems, law of large numbers
- Mathematical statistics: point estimation, confidence interval, hypothesis testing, *p*-value, significance level

#### **Teaching and learning method**

- Course in six blocks
- Class discussions, written homework (computer exercises, GRETL) submitted by groups of (3-5) students, presentations of homework by participants
- Final exam

#### Assessment of student work

- For grading, the written homework, presentation of homework in class, and a final written exam will be of relevance
- Weights: homework 40 %, final written exam 60 %
- Presentation of homework in class: students must be prepared to be called at random

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#### Organizational Issues

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#### Introduction to GRETL

## Empirical Economics Prior to 1930ies

The situation in the early 1930ies

- Theoretical economics aims at "operationally meaningful theorems";
   "operational" means purely logical mathematical deduction
- Economic theories or laws are seen as deterministic relations; no inference from data as part of economic analysis
- Ignorance of the stochastic nature of economic concepts
- Data: limited availability; time-series on agricultural commodities, foreign trade
- Use of statistical methods for
  - measuring theoretical coefficients, e.g., demand elasticities
  - representing business cycles

### Early Institutions

- Applied demand analysis: US Bureau of Agricultural Economics
- Statistical analysis of business cycles: H.L.Moore (Columbia University): Fourier periodogram; W.M.Persons et al. (Harvard): business cycle forecasting; US National Bureau of Economic Research (NBER)
- Cowles Commission for Research in Economics
  - Founded 1932 by A.Cowles: determinants of stock market prices?
  - Formalization of econometrics, development of econometric methodology
  - **R.Frisch, G.Tintner; European refugees**
  - J.Marschak (head 1943-55) recruited people like T.C.Koopmans, T.M.Haavelmo, T.W.Anderson, L.R.Klein
  - □ Interests shifted to theoretical and mathematical economics after 1950

#### Early Actors

- R.Frisch (Oslo Institute of Economic Research): econometric project, 1930-35; T.Haavelmo, O.Reiersol
- J.Tinbergen (Dutch Central Bureau of Statistics, Netherlands Economic Institute; League of Nations, Genova): macro-econometric model of Dutch economy, ~1935; T.C.Koopmans, H.Theil
- Austrian Institute for Trade Cycle Research (Österreichisches Institut für Konjunkturforschung, 1927, F.v.Hayek, L.v.Mises):
   O.Morgenstern (head), A.Wald, G.Tintner
- Econometric Society, founded 1930 by R.Frisch et al.
  - Facilitates exchange of scholars from Europe and US
  - Dealing with econometrics and mathematical statistics

#### First Steps

- R.Frisch, J.Tinbergen:
  - Macro-economic modelling based on time-series, ~ 1935
  - □ Aiming at measuring parameters, e.g., demand elasticities
  - Aware of problems due to quality of data
  - Nobel Memorial Prize in Economic Sciences jointly in 1969 ("for having developed and applied dynamic models for the analysis of economic processes")
- T.Haavelmo
  - "The Probability Approach in Econometrics": PhD thesis (1946)
  - Econometrics as a tool for testing economic theories
  - Nobel Memorial Prize in Economic Sciences in 1989 ("for his clarification of the probability theory foundations of econometrics and his analyses of simultaneous economic structures")

#### First Steps, cont'd

- Cowles Commission
  - Formalization of econometrics, development of econometric methodology
  - Methodology for macro-economic modelling based on Haavelmo's approach
  - Cowles Commission monographs by G.Tintner, T.C.Koopmans, et al.

### The Haavelmo Revolution

Introduction of probabilistic concepts in economics

- Obvious deficiencies of traditional approach: Residuals, measurement errors, omitted variables; stochastic time-series data
- Advances in probability theory in early 1930ies
- Fisher's likelihood function approach
- Haavelmo's ideas
  - Critical view of Tinbergen's macro-econometric models
  - Thorough adoption of probability theory in econometrics
  - Conversion of deterministic economic models into stochastic structural equation models
- Haavelmo's "The Probability Approach in Econometrics"
  - Why is the probability approach indispensable?
  - Modelling procedure based on ML estimation and hypothesis testing

### Cowles Commission Methodology

Assumptions based to macro-econometric modelling and testing of economic theories

Time series model

 $Y_t = \alpha X_t + \beta W_t + u_{1t}$ 

 $X_{\rm t} = \gamma Y_{\rm t} + \delta Z_{\rm t} + u_{\rm 2t}$ 

- 1. Specification of the model equation(s) includes the choice of variables; functional form is (approximately) linear
- 2. Time-invariant model equation(s): the model parameters  $\alpha$ , ...,  $\delta$  are independent of time *t*
- 3. Parameters  $\alpha$ , ...,  $\delta$  are structurally invariant, i.e., invariant wrt changes in the variables
- 4. Causal ordering (exogeneity, endogeneity) of variables is known
- 5. Statistical tests can falsify but not verify a model

### Classical Econometrics and More

- "Golden age" of econometrics until ~1970
  - Multi-equation models for analyses and forecasting
  - Growing computing power
  - Development of econometric tools
- Skepticism
  - Poor forecasting performance
  - Dubious results due to
    - wrong specifications
    - imperfect estimation methods
- Time-series econometrics: non-stationarity of economic time-series
  - Consequences of non-stationarity: misleading t-, DW-statistics, R<sup>2</sup>
  - Non-stationarity: needs new models (ARIMA, VAR, VEC); Box & Jenkins (1970: ARIMA-models), Granger & Newbold (1974, spurious regression), Dickey-Fuller (1979, unit-root tests)

Model	year	eq's
Tinbergen	1936	24
Klein	1950	6
Klein & Goldberger	1955	20
Brookings	1965	160
Brookings Mark II	1972	~200

#### Econometrics ...

- ... consists of the application of statistical data and techniques to mathematical formulations of economic theory. It serves to test the hypotheses of economic theory and to estimate the implied interrelationships. (Tinbergen, 1952)
- ... is the interaction of economic theory, observed data and statistical methods. It is the interaction of these three that makes econometrics interesting, challenging, and, perhaps, difficult. (Verbeek, 2008)
  - ... is a methodological science with the elements
    - economic theory
    - mathematical language
    - statistical methods
    - computer science

aiming to give empirical content to economic relations. (Pesaran, 1987)

#### Our Course

- 1. Introduction to linear regression (Verbeek, Ch. 2): the linear regression model, OLS method, properties of OLS estimators
- 2. Introduction to linear regression (Verbeek, Ch. 2): goodness of fit, hypotheses testing, multicollinearity
- 3. Interpreting and comparing regression models (Verbeek, Ch. 3): interpretation of the fitted model, selection of regressors, testing the functional form
- 4. Heteroskedascity and autocorrelation (Verbeek, Ch. 4): causes and consequences, testing, alternatives for inference
- 5. Endogeneity, instrumental variables and GMM (Verbeek, Ch. 5): the IV estimator, the generalized instrumental variables estimator, the generalized method of moments (GMM)
- 6. The practice of econometric modelling

### Econometrics 2: An Advanced Course

- Univariate and multivariate time series models: ARMA-, ARCH-, GARCH-models, VAR-, VEC-models
- Models for panel data
- Models with limited dependent variables: binary choice, count data

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  - MALE: gender (1 if male, 0 otherwise)
  - SCHOOL: years of schooling
  - EXPER: experience in years
  - AGE: age in years
- Possible questions
  - Effect of gender on wage p.h.: Average wage p.h.: 6,31\$ for males, 5,15\$ for females
  - □ Effects of education, of experience, of interactions, etc. on wage p.h.



#### Linear Regression

Y: explained variable
X: explanatory or regressor variable
The linear regression model describes the data-generating process of Y under the condition X

simple linear regression model

$$Y = \beta_1 + \beta_2 X$$

 $\beta_2$ : coefficient of *X* 

 $\beta_1$ : intercept

multiple linear regression model

$$Y = \beta_1 + \beta_2 X_2 + \ldots + \beta_K X_K$$

#### Fitting a Model to Data

Choice of values  $b_1$ ,  $b_2$  for model parameters  $\beta_1$ ,  $\beta_2$  of  $Y = \beta_1 + \beta_2 X$ , given the observations ( $y_i$ ,  $x_i$ ), i = 1,...,N

Principle of (Ordinary) Least Squares or OLS:  $b_i = \arg \min_{\beta_1, \beta_2} S(\beta_1, \beta_2), i = 1,2$ 

Objective function: sum of the squared deviations  $S(\beta_1, \beta_2) = \sum_i [y_i - (\beta_1 + \beta_2 x_i)]^2 = \sum_i \varepsilon_i^2$ 

Deviation between observation and fitted value:  $\varepsilon_i = y_i - (\beta_1 + \beta_2 x_i)$ 

### Observations and Fitted Regression Line

Simple linear regression: Fitted line and observation points (Verbeek, Figure 2.1)



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#### **OLS** Estimators

Equating the partial derivatives of  $S(\beta_1, \beta_2)$  to zero: normal equations

$$Nb_{1} + b_{2} \sum_{i=1}^{N} x_{i} = \sum_{i=1}^{N} y_{i}$$
$$b_{1} \sum_{i=1}^{N} x_{i} + b_{2} \sum_{i=1}^{N} x_{i}^{2} = \sum_{i=1}^{N} x_{i} y_{i}$$

OLS estimators  $b_1$  und  $b_2$  result in



with mean values  $\overline{x}$  and  $\overline{y}$ and second moments  $s_{xy} = \frac{1}{N-1} \sum_{i} (x_i - \overline{x})(y_i - \overline{y})$  $s_x^2 = \frac{1}{N-1} \sum_{i} (x_i - \overline{x})^2$ 

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### Individual Wages, cont'd

Sample (US National Longitudinal Survey, 1987): wage per hour, gender, experience, years of schooling; *N* = 3294 individuals (1569 females)

Average wage p.h.: 6,31\$ for males, 5,15\$ for females Model:

 $wage_i = \beta_1 + \beta_2 male_i + \varepsilon_i$ 

*male*<sub>I</sub>: male dummy, has value 1 if individual is male, otherwise value 0

**OLS** estimation gives

 $wage_{i} = 5,15 + 1,17^{*}male_{i}$ 

Compare with averages!



#### **OLS Estimators: General Case**

Model for Y contains K-1 explanatory variables

$$Y = \beta_1 + \beta_2 X_2 + \dots + \beta_K X_K = x'\beta$$

with  $x = (1, X_2, ..., X_K)'$  and  $\beta = (\beta_1, \beta_2, ..., \beta_K)'$ 

Observations:  $(y_i, x_i') = (y_i, (1, x_{i2}, ..., x_{iK})), i = 1, ..., N$ 

OLS estimates  $b = (b_1, b_2, ..., b_K)$ ' are obtained by minimizing the objective function wrt the  $\beta_k$ 's

$$S(\boldsymbol{\beta}) = \sum_{i=1}^{N} (y_i - x_i' \boldsymbol{\beta})^2$$

this results in

$$-2\sum_{i=1}^{N} x_i (y_i - x'_i b) = 0$$

### OLS Estimators: General Case,

#### cont'd

#### or

$$\left(\sum_{i=1}^{N} x_i x_i'\right) b = \sum_{i=1}^{N} x_i y_i$$

the **normal equations**, a system of *K* linear equations for the components of *b* 

Given that the symmetric *K*x*K*-matrix  $\sum_{i=1}^{N} x_i x'_i$  has full rank *K* and is hence invertible, the OLS estimators are

$$b = \left(\sum_{i=1}^{N} x_{i} x_{i}'\right)^{-1} \sum_{i=1}^{N} x_{i} y_{i}$$

#### **Best Linear Approximation**

Given the observations:  $(y_i, x_i') = (y_i, (1, x_{i2}, ..., x_{iK})), i = 1, ..., N$ 

For  $y_i$ , the linear combination or the fitted value

$$\hat{y}_i = x'_i b$$

is the best linear combination for Y from  $X_2, ..., X_K$  and a constant (the intercept)

#### Some Matrix Notation

N observations

$$(y_1, x_1), \ldots, (y_N, x_N)$$

Model: 
$$y_i = \beta_1 + \beta_2 x_i + \varepsilon_i$$
,  $i = 1, ..., N$ , or  
 $y = X\beta + \varepsilon$ 

with

$$y = \begin{pmatrix} y_1 \\ \vdots \\ y_N \end{pmatrix}, X = \begin{pmatrix} 1 & x_1 \\ \vdots & \vdots \\ 1 & x_N \end{pmatrix}, \beta = \begin{pmatrix} \beta_1 \\ \beta_2 \end{pmatrix}, \varepsilon = \begin{pmatrix} \varepsilon_1 \\ \vdots \\ \varepsilon_N \end{pmatrix}$$

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### OLS Estimators in Matrix Notation

Minimizing

 $S(\beta) = (y - X\beta)' (y - X\beta) = y'y - 2y'X\beta + \beta' X'X\beta$ 

with respect to  $\beta$  gives the normal equations

$$\frac{\partial S(\beta)}{\partial \beta} = -2(X'y - X'Xb) = 0$$

resulting from differentiating S( $\beta$ ) with respect to  $\beta$  and setting the first derivative to zero

The vector *b* of OLS solution or OLS estimators for  $\beta$  is

$$b = (XX)^{-1}XY$$

The best linear combinations or **predicted values** for *Y* given *X* or projections of *y* into the space of *X* are obtained as

 $\hat{y} = Xb = X(XX)^{-1}Xy = P_xy$ 

the NxN-matrix  $P_x$  is called the projection matrix or hat matrix

#### **Residuals in Matrix Notation**

The vector y can be written as  $y = Xb + e = \hat{y} + e$  with residuals

e = y - Xb or  $e_i = y_i - x_i'b$ , i = 1, ..., N

From the normal equations follows

 $-2(X^{c}y - X^{c}Xb) = -2 X^{c}e = 0$ 

i.e., each column of X is orthogonal to e

With

 $e = y - Xb = y - P_x y = (I - P_x)y = M_x y$ 

the **residual generating matrix**  $M_x$  is defined as

 $M_{\rm x} = I - X(XX)^{-1}X = I - P_{\rm x}$ 

 $M_x$  projects y into the orthogonal complement of the space of X

Properties of  $P_x$  and  $M_x$ : symmetry ( $P'_x = P_x$ ,  $M'_x = M_x$ ) idempotence ( $P_xP_x = P_x$ ,  $M_xM_x = M_x$ ), and orthogonality ( $P_xM_x = 0$ )
#### **Properties of Residuals**

Residuals:  $e_i = y_i - x_i$ , i = 1, ..., N

Minimum value of objective function

 $S(b) = e'e = \Sigma_i e_i^2$ 

From the orthogonality of  $e = (e_1, ..., e_N)$ ' to each  $x_i = (x_{1i}, ..., x_{Ni})$ ', i = 1, ..., K, i.e.,  $e'x_i = 0$ , follows that

 $\Sigma_i e_i = 0$ 

i.e., average residual is zero, if the model has an intercept

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#### US Wages

US wages are gender-specific

The relation

 $wage_i = \beta_1 + \beta_2 male_i + \varepsilon_i$ 

with *male*<sub>I</sub>: male dummy (equals 1 for males, otherwise 0)

- describes the wage of individual i as a function of its gender
- is assumed to be true for all US citizens

Given sample data (*wage*<sub>i</sub>, *male*<sub>i</sub>, *i* = 1,...*N*), OLS estimation of  $\beta_1$  and  $\beta_2$  may result in

 $wage_i = 5,15 + 1,17^*male_i$ 

- This is not (only) a description of the sample!
- But reflects a general relationship

#### Income and Consumption



Consumption function  $PCR_t = \beta_1 + \beta_2 PYR_t + \varepsilon_t$ • describes consumption in the Euro-zone



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#### **Economic Models**

Describe economic relationships (not just a set of observations), have an economic interpretation

Linear regression model:

$$y_i = \beta_1 + \beta_2 x_{i2} + \ldots + \beta_K x_{iK} + \varepsilon_i = x_i'\beta + \varepsilon_i$$

- Variables  $Y, X_2, ..., X_K$ : observable
- Error term ε<sub>i</sub> (disturbance term) contains all influences that are not included explicitly in the model; not observable; assumption E{ε<sub>i</sub> | x<sub>i</sub>} = 0 gives

$$\mathsf{E}\{y_i \mid x_i\} = x_i^{`}\beta$$

the model describes the expected value of y given x

- Sample  $(y_i, x_{i2}, ..., x_{iK}, i = 1, ..., N)$  from a well-defined population
- Unknown coefficients  $\beta_1, ..., \beta_K$ : population parameters

### Sampling in the Economic Context

The regression model  $y_i = x_i'\beta + \varepsilon_i$ , i = 1, ..., N; or  $y = X\beta + \varepsilon$ 

describes one realization out of all possible samples of size *N* from the population

A) Sampling process with fixed, i.e., non-stochastic  $x_i$ 's

- New sample: new error terms  $\varepsilon_i$ ,  $i = 1, ..., N_i$ , and, hence, new  $y_i$ 's
- Joint distribution of  $\varepsilon_i$ 's determines properties of *b* etc.
- A laboratory setting, does not apply to the economic context
- B) Sampling process with samples of  $(x_i, y_i)$  or  $(x_i, \varepsilon_i)$
- New sample: new error terms  $\varepsilon_i$  and new  $x_i$ , i = 1, ..., N
- Random sampling of  $(x_i, \varepsilon_i)$ , i = 1, ..., N: joint distribution of  $(x_i, \varepsilon_i)$ 's determines properties of *b* etc.

# Sampling in the Economic Context, cont'd

- The sampling with fixed, non-stochastic x<sub>i</sub>'s is not realistic for economic data
- Sampling process with samples of (x<sub>i</sub>, y<sub>i</sub>) is appropriate for modeling cross-sectional data
  - Example: household surveys, e.g., US National Longitudinal Survey, EU-SILC
- Sampling process with samples of (x<sub>i</sub>, y<sub>i</sub>) from time-series data: sample is seen as one out of all possible realizations of the underlying data-generating process
  - □ Example: time series PYR and PCR of the AWM-Database

### Assumptions of the Linear Regression Model

The linear regression model  $y_i = x_i'\beta + \varepsilon_i$  makes use of assumptions

- Assumption for  $\varepsilon_i$ 's: E{ $\varepsilon_i | x_i$ } = 0; exogeneity of variables X
  - X contains no information on the error term  $\varepsilon$
  - $E\{\varepsilon_i \mid x_i\} = 0$  implies that  $\varepsilon_i$  and  $x_i$  are uncorrelated
- This implies

 $\mathsf{E}\{y_i \mid x_i\} = x_i'\beta$ 

i.e., the regression line describes the conditional expectation of  $y_i$  given  $x_i$ 

Coefficient β<sub>k</sub> measures the change of the expected value of Y if X<sub>k</sub> changes by one unit and all other X<sub>j</sub> values, j + k, remain the same (ceteris paribus condition)

#### **Regression Coefficients**

Linear regression model:

$$y_i = \beta_1 + \beta_2 x_{i2} + \ldots + \beta_K x_{iK} + \varepsilon_i = x_i'\beta + \varepsilon_i$$

Coefficient  $\beta_k$  measures the change of the expected value of Y if  $X_k$  changes by one unit and all other  $X_j$  values,  $j \neq k$ , remain the same (ceteris paribus condition); marginal effect of changing  $X_k$  on Y

$$\frac{\partial E\{y_i | x_i\}}{\partial x_{ik}} = \beta_k$$

Example

• Wage equation:  $wage_i = \beta_1 + \beta_2 male_i + \beta_3 school_i + \beta_4 exper_i + \varepsilon_i$ 

 $\beta_3$  measures the impact of one additional year at school upon a person's wage, keeping gender and years of experience fixed

#### Estimation of β

Given a sample  $(x_i, y_i)$ , i = 1, ..., N, the OLS estimators for  $\beta$  $b = (XX)^{-1}Xy$ 

can be used as an approximation for  $\beta$ 

- The vector *b* is a vector of numbers, the estimates
- The vector b is the realization of a vector of random variables
- The sampling concept and assumptions on  $\varepsilon_i$ 's determine the quality, i.e., the statistical properties, of *b*

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### Fitting Economic Models to Data

**Observations allow** 

- to estimate parameters
- to assess how well the data-generating process is represented by the model, i.e., how well the model coincides with reality
- to improve the model if necessary

Fitting a linear regression model to data provides

- Parameter estimates  $b = (b_1, ..., b_K)$ ' for coefficients  $\beta = (\beta_1, ..., \beta_K)$ '
- standard errors  $se(b_k)$  of the estimates  $b_k$ , k=1,...,K
- *t*-statistics, *F*-statistic, *R*<sup>2</sup>, Durbin Watson test-statistic, etc.

#### Individual Wages, cont'd

Wage equation with three regressors (Table 2.2, Verbeek)

**Table 2.2**OLS results wage equation

Dependent variable: wage

Variable	Estimate	Standard error	r <i>t</i> -ratio
constant <i>male</i> school exper	-3.3800 1.3444 0.6388 0.1248	$\begin{array}{c} 0.4650 \\ 0.1077 \\ 0.0328 \\ 0.0238 \end{array}$	-7.2692 12.4853 19.4780 5.2530
s = 3.0462	$R^2 = 0.1326$	$\overline{R}^2 = 0.1318$	F = 167.63

### OLS Estimator and OLS Estimates *b*

OLS estimates *b* are a realization of the OLS estimator

The OLS estimator is a random variable

- Observations are a random sample from the population
- Observations are generated by some random sampling process
  Distribution of the OLS estimator
- Actual distribution not known
- Distribution determined by assumptions on
  - model specification
  - the error term  $\varepsilon_i$  and regressor variables  $x_i$

Quality criteria (bias, accuracy, efficiency) of OLS estimates are determined by the properties of this distribution

#### Gauss-Markov Assumptions

Observation  $y_i$  is a linear function

$$y_i = x_i'\beta + \varepsilon_i$$

of observations  $x_{ik}$  of the regressor variables  $X_k$ , k = 1, ..., K, and the error term  $\varepsilon_i$ 

for 
$$i = 1, ..., N$$
;  $x_i' = (x_{i1}, ..., x_{iK})$ ;  $X = (x_{ik})$ 

A1	$E\{\varepsilon_i\} = 0$ for all <i>i</i>
A2	all $\varepsilon_i$ are independent of all $x_i$ (exogenous $x_i$ )
A3	$V{\varepsilon_i} = \sigma^2$ for all <i>i</i> (homoskedasticity)
A4	Cov{ $\varepsilon_i$ , $\varepsilon_j$ } = 0 for all <i>i</i> and <i>j</i> with $i \neq j$ (no autocorrelation)

In matrix notation:  $E{\epsilon} = 0$ ,  $V{\epsilon} = \sigma^2 I_N$ 

#### Systematic Part of the Model

The systematic part  $E\{y_i | x_i\}$  of the model  $y_i = x_i'\beta + \varepsilon_i$ , given observations  $x_i$ , is derived under the Gauss-Markov assumptions as follows:

(A2) implies  $E\{\varepsilon \mid X\} = E\{\varepsilon\} = 0$  and  $V\{\varepsilon \mid X\} = V\{\varepsilon\} = \sigma^2 I_N$ 

- Observations  $x_i$ , i = 1, ..., N, do not affect the properties of  $\varepsilon$
- The systematic part

 $\mathsf{E}\{y_i \mid x_i\} = x_i'\beta$ 

can be interpreted as the conditional expectation of  $y_i$ , given observations  $x_i$ 

## Is the OLS Estimator a Good Estimator?

- Under the Gauss-Markov assumptions, the OLS estimator has favourable properties; see below
- Gauss-Markov assumptions are very strong but not always satisfied
- Relaxations of the Gauss-Markov assumptions and consequences of such relaxations are important topics in econometrics

#### **Properties of OLS Estimators**

1. The OLS estimator *b* is unbiased:  $E\{b \mid X\} = E\{b\} = \beta$ Needs assumptions (A1) and (A2)

2. The variance of the OLS estimator *b* is given by

 $V\{b \mid X\} = V\{b\} = \sigma^2(\Sigma_i x_i x_i')^{-1} = \sigma^2(X' X)^{-1}$ 

Needs assumptions (A1), (A2), (A3) and (A4)

3. Gauss-Markov Theorem: The OLS estimator b is a BLUE<sup>1)</sup> (best linear unbiased estimator) for β
 Needs assumptions (A1), (A2), (A3), and (A4) and requires linearity in parameters

<sup>&</sup>lt;sup>1)</sup> OLS estimator is most accurate among linear unbiased estimators; see next slide

#### The Gauss-Markov Theorem

OLS estimator *b* is BLUE (best linear unbiased estimator) for  $\beta$ 

- Linear estimator: b\* = Ay with any full-rank KxN matrix A
- $b^*$  is an unbiased estimator:  $E\{b^*\} = E\{Ay\} = \beta$
- b is BLUE: V{b\*} V{b} is positive semi-definite, i.e., the variance of any linear combination d'b\* is not smaller than that of d'b
  V{d'b\*} ≥ V{d'b}

e.g.,  $V{b_k^*} \ge V{b_k}$  for any k

 The OLS estimator is most accurate among the linear unbiased estimators

### Standard Errors of OLS Estimators

Variance (covariance matrix) of the OLS estimators:

 $V\{b\} = \sigma^{2}(X X)^{-1} = \sigma^{2}(\Sigma_{i} x_{i} x_{i}')^{-1}$ 

- Standard error of OLS estimate b<sub>k</sub>: The square root of the k<sup>th</sup> diagonal element of V{b}
- V{*b*} is proportional to the variance  $\sigma^2$  of the error terms
- Estimator for  $\sigma^2$ : sampling variance  $s^2$  of the residuals  $e_i$

 $s^2 = (N - K)^{-1} \Sigma_i e_i^2$ 

Under assumptions (A1)-(A4),  $s^2$  is unbiased for  $\sigma^2$ 

Attention: the estimator  $(N - 1)^{-1} \Sigma_i e_i^2$  is biased

Estimated variance (covariance matrix) of *b*:

 $\tilde{V}{b} = s^2(X X)^{-1} = s^2(\Sigma_i x_i x_i)^{-1}$ 

# Estimated Standard Errors of OLS Estimators

Variance (covariance matrix) of the OLS estimators:

 $V\{b\} = \sigma^{2}(X X)^{-1} = \sigma^{2}(\Sigma_{i} x_{i} x_{i}')^{-1}$ 

Standard error of OLS estimate b<sub>k</sub>: The square root of the k<sup>th</sup> diagonal element of V{b}

σ√c<sub>kk</sub>

with  $c_{kk}$  the k-th diagonal element of  $(X X)^{-1}$ 

Estimated variance (covariance matrix) of b:

 $\tilde{V}{b} = s^2(X X)^{-1} = s^2(\Sigma_i x_i x_i)^{-1}$ 

Estimated standard error of b<sub>k</sub>:

$$se(b_k) = s\sqrt{c_{kk}}$$

#### Two Examples

1. Simple regression  $y_i = \alpha + \beta x_i + \varepsilon_t$ 

The variance for the OLS estimator of  $\beta$  is

$$V\{b\} = \frac{\sigma^2}{Ns_x^2}$$

*b* is the more accurate, the larger *N* and  $s_x^2$  and the smaller  $\sigma^2$ 

2. Regression with two regressors:

 $y_{i} = \beta_{1} + \beta_{2} x_{i2} + \beta_{3} x_{i3} + \varepsilon_{t}$ 

The variance for the OLS estimator of  $\beta_2$  is

$$V\{b_2\} = \frac{1}{1 - r_{23}^2} \frac{\sigma^2}{Ns_{x2}^2}$$

 $r_{23}^2$ : correlation coefficient between  $X_2$  and  $X_3$  $b_2$  is most accurate if  $X_2$  and  $X_3$  are uncorrelated

### Normality of Error Terms

For the purpose of statistical inference, a distributional assumption for the  $\varepsilon_i$ 's is needed

A5  $\varepsilon_i$  normally distributed for all *i* 

Together with assumptions (A1), (A3), and (A4), (A5) implies

 $\varepsilon_i \sim \text{NID}(0, \sigma^2)$  for all *i* 

i.e., all  $\varepsilon_i$  are

- independent drawings
- from the *normal* distribution
- with mean 0
- and variance  $\sigma^2$

Error terms are "normally and independently distributed" (NID)

#### **Properties of OLS Estimators**

1. The OLS estimator *b* is unbiased:  $E\{b\} = \beta$ 

2. The variance of the OLS estimator is given by

 $V\{b\} = \sigma^2(X'X)^{-1}$ 

- 3. The OLS estimator b is a BLUE (best linear unbiased estimator) for  $\beta$
- 4. The OLS estimator *b* is normally distributed with mean  $\beta$  and covariance matrix V{*b*} =  $\sigma^2(X^tX)^{-1}$

 $b \sim N(\beta, \sigma^2(XX)^{-1}), b_k \sim N(\beta_k, \sigma^2 c_{kk})$ with  $c_{kk}$ : the *k*-th diagonal element of  $(XX)^{-1}$ Needs assumptions (A1) - (A5)

# Individual Wages: Relevance of Assumptions

 $wage_i = \beta_1 + \beta_2^* male_i + \varepsilon_i$ 

What do the assumptions mean?

- (A1):  $\beta_1 + \beta_2^* male_i$  contains the entire systematic part of the model; no other regressors besides gender are relevant?
- (A2):  $x_i$  uncorrelated with  $\varepsilon_i$  for all *i*: knowledge of a person's gender provides no information about further variables which affect the person's wage; is this realistic?
- (A3) V{ $\epsilon_i$ } =  $\sigma^2$  for all *i*: variance of error terms (and of wages) is the same for males and females; is this realistic?

(A4) Cov{ $\varepsilon_{i,}, \varepsilon_{j}$ } = 0,  $i \neq j$ : implied by random sampling

(A5) Normality of  $\varepsilon_i$ : is this realistic? (Would allow, e.g., for negative wages)



#### Your Homework

- Verbeek's data set "wages1" contains for a sample of 3294 individuals the wage p.h. (wage) and other variables. Calculate, using GRETL, for the variable school (years of schooling) the mean (a) of the whole sample, (b) of males and females, and (c) the standard deviation of the years of schooling for males and for females.
- 2. For Verbeek's data set "wages1", using GRETL, (a) cross-tabulate the variable *school* (years of schooling) over *male* for individuals with *school* at least 8 years; compare (b) the mean values of the males and females; draw for the whole population (c) scatter plots of *wage* over *school* and *exper*; and (d) a factorized box plot of *wage* over *school*. Discuss the results.

#### Your Homework, cont'd

- 3. For the simple regression  $y_i = \alpha + \beta x_i + \varepsilon_i$ , i = 1,...,N, show that the variance of the OLS estimate for  $\beta$  is  $\sigma^2/(Ns_x^2)$ , where  $\sigma^2$  is the error term variance,  $s_x^2$  the variance of the  $x_i$ 's.
- 4. For the sample  $(y_i, x_i)$ , i = 1, ..., N, and the linear regression  $(y_i = \beta_1 + \beta_2 x_i + \varepsilon_i)$ : (a) write out the matrices XX and Xy; (b) write out the determinant  $det[(XX)^{-1}]$ , the matrix  $(XX)^{-1}$ , and the OLS estimator  $b = (XX)^{-1}Xy$ .