# M7777 Applied Functional Data Analysis 9. Functional Response with Scalar Covariate

#### Jan Koláček (kolacek@math.muni.cz)

Dept. of Mathematics and Statistics, Faculty of Science, Masaryk University, Brno



#### **Functional ANOVA**

Just as in the standard ANOVA, let be K ( $K \ge 3$ ) groups and

$$x_{ij}(t) \dots i$$
-th curve in *j*-th group,  $i = 1, \dots, n_j, \ n = \sum_{j=1}^{n} n_j$ 

An over-all mean

$$\hat{\mu}(t) = \bar{x}(t) = \frac{1}{n} \sum_{j=1}^{K} \sum_{i=1}^{n_j} x_{ij}(t)$$

• Effects for each group

$$\hat{\alpha}_j(t) = \frac{1}{n_j} \sum_{i=1}^{n_j} (x_{ij}(t) - \bar{x}(t))$$

An error process

$$\hat{\varepsilon}_{ij}(t) = x_{ij}(t) - \hat{\alpha}_j(t) - \bar{x}(t)$$

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#### **Functional ANOVA model**

$$x_{ij}(t) = \mu(t) + \alpha_j(t) + \varepsilon_{ij}(t)$$
(1)

Suppose we observe scalar covariates  $z_{i1}, \ldots, z_{iK}$ 

$$z_{ij} = \begin{cases} 1 & \text{if } x_{ij}(t) \text{ belongs to } j\text{-th group} \\ 0 & \text{otherwise} \end{cases}$$

Model (1) can be rewritten as a functional linear model

$$y_i(t) = \beta_0(t) + \sum_{j=1}^{K} \beta_j(t) z_{ij} + \varepsilon_i(t)$$

with conditions 
$$\sum_{j=1}^{K} \beta_j(t) = 0$$
,  $E\varepsilon_i(t) = 0$ .

#### **Canadian Weather**

Divide all 35 locations to 4 regions: Atlantic, Continental, Pacific, Arctic



We will study the effect of geographic region on the shape of the temperature curves.

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Let us denote model parameters

 $\beta_0(t) \dots$  Canada  $\beta_1(t) \dots$  Atlantic  $\beta_2(t) \dots$  Continental  $\beta_3(t) \dots$  Pacific  $\beta_4(t) \dots$  Arctic

Estimates of  $\beta_j(t)$ 



Prediction of the temperature curve in j-th group

$$\hat{y}_j(t) = \hat{\beta}_0(t) + \hat{\beta}_j(t), \qquad j = 1, \dots, 4.$$



#### **Confidence Intervals**

Consider a basis representation

$$\beta_j(t) = \mathbf{\Phi}_j(t) \mathbf{c}_j$$

set  $\mathbf{b} = (\mathbf{c}'_0, \dots, \mathbf{c}'_K)' \Rightarrow \hat{\beta}_j(t)$  depends on  $\hat{\mathbf{b}}$ . Let  $\mathbf{t} = (t_1, \dots, t_N)$ ,  $\mathbf{y}_i = (y_i(t_1), \dots, y_i(t_N))'$ , generally, we minimize penalized least squares and get the estimate

$$\hat{\mathbf{b}} = y2cMap\,\mathbf{y}.$$

Estimate the covariance matrix

$$\hat{\boldsymbol{\Sigma}} = rac{1}{n-\mathcal{K}}\sum_{i=1}^{n} \hat{arepsilon}_{i}\hat{arepsilon}_{i}^{\prime}, \quad ext{where } \hat{arepsilon}_{i} = \mathbf{y}_{i} - \mathbf{z}_{i}\hat{eta}(\mathbf{t}).$$

Thus (formally)

$$\operatorname{Var}\hat{\beta}_{j}(t) = \mathbf{\Phi}_{j}(t)$$
y2cMap $\hat{\mathbf{\Sigma}}$ y2cMap $\mathbf{\Phi}_{j}(t)'$ .

Estimates of  $\beta_j(t)$ 





#### Assessing the fit of the fANOVA



Functional  $R^2$ 





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#### F-statistic

To test significance, we can define a pointwise F-statistic

$$F(t) = \frac{\operatorname{Var}(\hat{\mathbf{y}}(t))}{\sum\limits_{i=1}^{n} (y_i(t) - \hat{y}_i(t))^2 / n}$$

indicates where there is a large amount of signal relative to variance.

Test over-all regression significance based on

$$F^* = \max F(t).$$

#### **Permutation Test**

We would like to test the null hypothesis

$$H_0: Ey(t) = 0 \quad \forall t \in [t_1, t_N]$$

Do B times

- Permute indexes  $1, \ldots, n$  to get  $i_1, \ldots, i_n$ , leaving the design unchanged.
- **2** Define  $y_j^b(t) = y_{i_j}(t)$ .
- **3** Estimate the model using  $\mathbf{y}^{b}(t)$  as the response.
- $\textbf{\textbf{\textbf{M}} easure } F_b^* \text{ and set } I_b = \begin{cases} 1 & \text{if } F_b^* > F^* \\ 0 & \text{if } F_b^* \le F^* \end{cases}$

Then *p*-value for the test

$$p_B = rac{1}{B}\sum_{b=1}^B I_b$$

#### **Canadian Weather**

 $B = 200, p_B = 0$ 



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#### **Canadian Weather**

detailed test results



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#### **Canadian Weather**





200

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50

0

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150

100

Sample

#### Functional *t*-test

Just 2 groups of curves  $(x_{ij}(t), x_{i2}(t))$ : Is the difference statistically significant?

#### Example. Berkeley Growth Study (39 boys, 54 girls)



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#### Functional *t*-statistic

To test significance, we can define a pointwise *t*-statistic

$$T(t) = \frac{|\bar{x}_1(t) - \bar{x}_2(t)|}{\sqrt{\frac{1}{n_1} \mathsf{Var}[x_1(t)] + \frac{1}{n_2} \mathsf{Var}[x_2(t)]}}$$

indicates where there is a large mean difference relative to variance.

Test over-all significance based on

$$T^* = \max T(t).$$

#### **Permutation Test**

We would like to test the null hypothesis

$$H_0: Ex_1(t) = Ex_2(t) \quad \forall t \in [t_1, t_N]$$

Do B times

- 1 Randomly shuffle the labels of the curves.
- 2 Calculate the *t*-statistic  $T_b(t)$  with the new labels.

$$\textbf{3} \text{ Measure } T_b^* \text{ and set } I_b = \begin{cases} 1 & \text{ if } T_b^* > T^* \\ 0 & \text{ if } T_b^* \le T^* \end{cases}$$

Then *p*-value for the test

$$p_B = \frac{1}{B} \sum_{b=1}^{B} I_b$$

#### Berkeley Growth Study

 $B = 200, p_B = 0$ 



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### Berkeley Growth Study

#### detailed test results



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#### **Berkeley Growth Study**



#### Sound Intensity Data

Load the variable rat3 from the rat3.RData file. The variable rat3 contains observations of a rat neural activity evoked by sound intensity. The evoked potential (EPI) was measured in dependence on 19 sound intensities for 5 days. The dataset contains 79 repetitions for each day.

- Smooth the data by B-spline bases with second-derivative penalties and plot the result with color-day specification (see Figure 1).
- Conduct a study of the effect of the day on the shape of the EPI curves. Consider the fANOVA model with days as covariates. Plot estimated parameters with its pointwise confidence bands (see Figure 2).
- Plot predictions for each day with its pointwise confidence bands (see Figure 3).
- Plot functional  $R^2$  of the model (see Figure 4) and interpret it.
- Asses the model by the permutation test for *F*-statistic and plot the result (see Figure 5).

- 2 Sound Intensity Data
  - Consider just days SS4 and SS5 and plot the EPI estimates with color-day specification (see Figure 6).
  - Is the difference between days statistically significant? Conduct the functional *t*-test.
  - Asses the model by the permutation test for *t*-statistic, plot the result (see Figure 7) and interpret it.





Figure 2.



Figure 3.



Figure 4.

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Figure 6.

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